

PROOF OF THE LAWS OF TWIN-BIRTHS.

By SIR GEORGE H. KNIBBS, C.M.G., M.Inst.Int.d.Stat.,
Hon.F.S.S., Hon.M.Amer.S.A., Hon.M.Stat.Soc.Paris, and of
Hungary, etc.

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SYNOPSIS. 1. Formulæ giving separately the relative numbers of uniovular and dioivular twins. 2. Applicability to total twins. 3. Applicability to cases of two males, male and female, and two females. Limitations of the formulæ.

1. *Formulae giving separately the relative numbers of uniovular and dioivular twins.* In an earlier paper, this Journal, Vol. 61, pp. 190-217, I showed that the frequency of twins born through the subdivision of a single fertilised ovum was sensibly the same for each age of the mothers, being always about 0.00300 of the cases of maternity. I showed, also, that the numbers born from two ova increased sensibly linearly up to age, say, 37 or 38, then decreased nearly linearly to the end of the child-bearing period (see formulae (2) and (2a), p. 200; and (3) and (3a), p. 202, this Journal). In formulae (3) the sign should be +. From this it is evident that the numbers for each age for the two types of cases can be given by the following type of formulae, in which t denotes the ratio of the number of twins to the number of cases of maternity:—

Uniovular cases. + Dioivular cases.

Up to age 37	$t=0.00300$	+0.00058 ($x-15$)
Beyond age 37	$t=0.00300$	+0.01276—0.00128 ($x-37$)

The constants of these formulae have been computed to fit approximately with few figures, the Australian results for the six years 1920 to 1925. The number for each age

of cases of maternity (mothers) on which these are based are given in Table VI., column (ii), pp. 197-8, of this Journal.

2. *Applicability to total twins.* In order to obtain a definitive idea of how these expressions do actually represent the data, the values of the two parts of Ct are given as against the data. In making the comparisons, however, it has to be remembered that the ages are merely *as stated* by the mothers (see remarks p. 191, this Journal). The actual ages probably differ slightly, for the ages are not given exactly.

Between the ages 12 and 49 inclusive there were 809,126 mothers (cases of maternity) and 8,403 cases of twins, viz., MM 2,792, MF 2,989, and FF 2,622. This would indicate that the probable number of uniovular cases was $2,792 + 2,622 - 2,989$, that is 2,425. The calculated sum with the adopted approximate coefficient 0.00300, was 2,427.3.

It has already been shown (pp. 211-2 of this Journal) that this ratio is the same for each age during the child-bearing period.

Having regard to the number of cases, the rough formulae, given above, are seen from the table to represent closely the laws enunciated in the beginning of this paper, inasmuch as the observed and calculated results differ from each other for any one age, never more than -32 and $+30$ (ages 38 and 40 respectively), and the differences of the sums of these from age 12 upwards never disagree by more than -58 to $+23$ (ages 39 and 19 respectively).

The paucity of the numbers of cases suggests that it is not worth while attempting at present to furnish formulae of higher numerical precision.

The results thus established are:—(a) *three twins are produced, each from a single ovum, in every thousand cases*

Table I. Showing calculated and observed numbers of twins, for each age of mother from 13 to 49, and the sums of their differences.

Ages.	Uniov.	Diov.	Calc.	Obsd.	Diff.	Sum of Diff.
i	ii	iii	iv	v	vi	vii
13	0.1	0.0	0	1	+ 1	+ 1
14	0.5	0.0	1	1	0	+ 1
15	1.5	0.0	2	3	+ 1	+ 2
16	6.3	1.2	7	11	+ 4	+ 6
17	17.1	6.6	24	22	— 2	+ 4
18	35.2	20.4	56	54	— 2	+ 2
19	56.4	43.6	100	121	+21	+23
20	75.2	72.7	148	134	—14	+ 9
21	96.6	112.1	209	205	— 4	+ 5
22	114.1	154.5	270	266	— 4	+ 1
23	129.5	200.3	330	337	+ 7	+ 8
24	138.5	240.9	379	349	—30	—22
25	142.2	274.9	417	437	+20	— 2
26	144.6	307.5	452	443	— 9	—11
27	141.9	329.2	471	469	— 2	—13
28	141.3	355.2	497	466	—31	—44
29	135.7	367.3	503	521	+18	—26
30	137.8	399.5	537	545	+ 8	—18
31	114.2	353.2	467	458	— 9	—27
32	116.0	381.3	497	486	—11	—38
33	103.2	359.0	462	483	+21	—17
34	96.2	353.3	450	456	+ 6	—11
35	86.9	336.1	423	419	— 4	—15
36	79.4	322.4	402	403	+ 1	—14
37	68.1	289.7	358	362	+ 4	—10
38	62.8	240.2	303	271	—32	—42
39	52.0	176.7	229	213	—16	—58
40	42.1	125.2	167	197	+30	—28
41	29.1	74.1	103	98	— 5	—33
42	25.0	52.9	78	84	+ 6	—27
43	16.9	28.7	46	41	— 5	—32
44	10.1	12.8	23	29	+ 6	—26
45	6.0	5.1	11	11	0	—26
46	2.7	1.1	4	6	+ 2	—24
47	1.3	—	1	0	— 1	—25
48	0.6	—	1	1	0	—25
49	0.2	—	0	0	0	—25

Column (iv) is the sum of columns (ii) and (iii) cast to the nearest unit. Column (vi) is column (v) less column (iv). Column (vii) is the algebraic sums of the figures in column (vi).

of maternity, whatever the age of the mother; and (b) two ova are effectively fertilised (i.e., result in living twins), with a frequency increasing linearly with age till, say, age 37, but afterwards diminishing linearly till, say, about age 49 (or the end of the child-bearing period).

3. *Applicability of formulae to cases of two males, a male and a female, and two females.* In order to divide the calculated number of twins into MM, MF, and FF cases, it is necessary to know how the masculinity over all varies, if at all, with the age of the mother, and how the masculinity in uniovular cases compares with that of dioivular cases.

It has already been shown in this Journal, sections 8 and 9 of former paper, p. 202, that the changes in masculinity are complex and irregular. Its general trend over all (m-f)/(m-f), see Table IX., may be approximately indicated by the following linear equations, viz:—

$$\text{Ages 17 to 30; } \mu' = -0.0330 + 0.0074 (x - 17).$$

$$\text{Ages 30 to 38; } \mu' = +0.0632 - 0.0104 (x - 30).$$

$$\text{Ages 38 to 46; } \mu' = -0.0200 + 0.0046 (x - 38).$$

The masculinity over all, however, is 0.02023, and is thus not very different from that of first-births, and of all births, since the former for 1921 to 1925 was 0.0278, and that for all births was 0.0250 (see Table IV., p. 195, and also column (viii.), Table VI., p. 197, this Journal).

While it will not be possible to compute exactly the MM, MF, and FF cases for each age, they can be closely approximated on the assumption that the masculinity over all, was for each age $\lambda = 0.02023$, and that the masculinity was $2\lambda = 0.04046$ for the MM and FF dioivular cases, the former applying to the uniovular cases (see p. 203 of this Journal).

On the assumptions indicated we obtain from columns (ii.) and (iii.) in the preceding Table, the numbers of MM, MF, and FF cases. The respective results are:—

$$\begin{array}{llll}
 U; & \text{Uniovular cases, age } x, & \frac{1}{2}U_x(1+\lambda) & \frac{1}{2}U_x(1+\lambda) \\
 D; & \text{Diovular cases, age } x, & \frac{1}{4}D_x(1+2\lambda) & \frac{1}{2}D_x \quad \frac{1}{4}D_x(1+2\lambda) \\
 U+D; & \text{Total cases, age } x. & \text{MM:} & \text{MF; FF.}
 \end{array}$$

For example for age 37 we have from the tabular results, 68.1 and 289.7 in Table I., the following:—

		MM	MF	FF	Totals
Masc.* 0.0202;	Uniovular	34.7	0.0	33.4	68.1
Masc.* 0.0405;	Diovular	75.3	144.9	69.5	289.7
Masc. 0.0202;	Totals	110.0	144.9	102.9	357.8
According to the data ...		105	148	109	362
Difference (obsd.-calc.) ..		—5	+3	+6	+4

The masculinities marked * are for the MM and FF cases alone. The discrepancies —5, +3 and +6 are not large considering the paucity in the number of cases.

Table II. below gives the calculated results for each age to the nearest whole number, and also the observed results, for the MM, MF, and FF cases; it discloses, therefore, how far the formulae, which are general, agree with the data. It will be seen that it establishes the fact that (c) *over the entire child-bearing period*, excepting, perhaps, say the first three or four commencing and closing years thereof, *the uniovular cases for each year of age are equal to the sum of the cases of two males and two females, less the cases of male and female.*

The measure of the goodness of fit of the formulae as representative of the data is not merely the closeness of the MM, MF, and FF individual results for each age, but also the aggregate of such results, viz., the sums of MM, MF, and FF up to each age. It is evident from the table giving these that the dictum is confirmed, viz., that the number of diovular cases increases sensibly linearly to age 37, and diminishes very nearly linearly afterwards. This is clearly the general law.

Table II. Showing Calculated and Observed Distribution of Twins, Australia, 1920 to 1925, for each age and their sums up to each age.

Mother's	Ca.	Ob.	Ca.	Ob.	Ca.	Ob.	Ca.	Ob.	Ca.	Ob.	Ca.	Ob.
Age.	MM	MM	MF	MF	FF	FF	Sum	Sum	Sum	Sum	Sum	Sum
	MM	MM	MF	MF	FF	FF	MM	MM	MF	MF	FF	FF
13 ..	0	0	0	0	0	1	0	0	0	0	0	1
14 ..	1	0	0	0	0	1	1	0	0	0	0	2
15 ..	1	1	0	0	1	2	2	1	0	0	1	4
16 ..	4	4	0	1	3	6	6	5	0	1	4	10
17 ..	11	8	3	4	10	10	17	13	3	5	14	20
18 ..	23	25	10	7	22	22	40	38	13	12	36	42
19 ..	40	39	22	33	38	49	80	77	35	45	74	91
20 ..	57	56	36	36	54	42	137	133	71	81	128	133
21 ..	78	73	56	42	74	90	215	206	127	123	202	223
22 ..	98	95	77	73	93	98	313	301	204	196	295	321
23 ..	118	120	100	99	112	118	431	421	304	295	407	439
24 ..	133	127	120	108	126	114	564	548	424	403	533	553
25 ..	144	149	137	144	136	144	708	697	561	547	669	697
26 ..	154	152	154	152	145	139	862	849	715	699	814	836
27 ..	158	152	165	166	149	151	1020	1001	880	865	963	987
28 ..	164	156	178	171	154	139	1184	1157	1058	1036	1117	1126
29 ..	165	171	184	192	155	158	1349	1328	1242	1228	1272	1284
30 ..	174	197	200	209	163	139	1523	1525	1442	1437	1405	1423
31 ..	150	145	177	170	141	143	1673	1670	1619	1607	1546	1566
32 ..	158	166	191	195	148	125	1819	1836	1810	1802	1694	1691
33 ..	146	154	179	185	137	144	1965	1990	1989	1987	1831	1835
34 ..	141	148	177	164	132	144	2106	2138	2166	2151	1963	1979
35 ..	132	128	168	172	123	119	2238	2266	2334	2323	2086	2098
36 ..	124	118	161	162	116	123	2362	2384	2495	2485	2202	2221
37 ..	110	105	145	148	103	109	2472	2489	2640	2633	2305	2330
38 ..	95	95	120	89	88	87	2567	2584	2760	2722	2393	2417
39 ..	72	62	88	97	68	54	2639	2646	2848	2819	2461	2471
40 ..	54	58	63	73	51	66	2693	2704	2911	2892	2512	2537
41 ..	34	33	37	33	32	32	2727	2737	2948	2925	2544	2569
42 ..	27	28	26	33	25	23	2754	2765	2974	2958	2569	2592
43 ..	16	11	14	18	15	12	2770	2776	2988	2976	2584	2604
44 ..	9	11	6	8	8	10	2779	2787	2994	2984	2592	2614
45 ..	4	3	3	5	2	3	2783	2790	2997	2989	2594	2617
46 ..	2	2	1	0	1	4	2785	2792	2998	2989	2595	2621
47 ..	1	0	0	0	0	0	2786	2792	2998	2989	2595	2621
48 ..	1	0	0	0	0	1	2787	2792	2998	2989	2595	2622
49 ..	0	0	0	0	0	0	2787	2792	2998	2989	2595	2622

Why there should be this sudden change after age 37, in the frequency of the effective fertilisation of two ova, is not apparent. The changes in the frequency of mater-

nity, column (ii.), pp. 197-8, of this Journal, throw no light upon it; and so far as the writer is aware, there are no other related statistical progressions which give any indication of a possible cause.

4. *Limitations of the formulae.* Obviously, it is practically certain that the number of uniovular cases is not represented by $0.00300C$ at the very beginning and end of the child-bearing period, although otherwise it is sensibly constant throughout. This needs investigation with a much larger number of cases. It also seems probable that the straight lines, representing the changes of frequency with age, are terminated by short curves concave upwards, and also that the two lines are joined by a curve convex upwards. To determine whether these are so or not, also requires as data a far larger number of cases. So far as I am aware, the necessary data are not yet available. It is to be hoped in future the ages of the mothers can be more accurately ascertained. An examination of the data will show that the numbers attributed to certain ages should probably be assigned to adjoining ages.

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